Introduction to the Theory of Calibration

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- Reminder on the theory of sampling
- Regression estimator
- Calibration estimators, choice of the function of calibration
- General remarks
- What is a good calibration software?

- Finite population $U = \{1, 2, ..., k, ..., N\}$
- Variable of interest y.
- Values taken by the variable of interest on the population

 $(y_1,\ldots,y_k,\ldots,y_N).$

Functions of interest

• Total
$$Y = \sum_{k \in U} y_k$$

• Mean
$$\overline{Y} = \frac{1}{N} \sum_{k \in U} y_k$$

• Variance
$$S_y^2 = \frac{1}{N} \sum_{k \in U} (y_k - \overline{Y})^2$$

• A sample s is a subset of the population U.



 A sampling design is a probability distribution on all the possible samples:

$$p(s) \ge 0$$
, for all $s \subset U$, and $\sum_{s \subset U} p(s) = 1$.

• The random sample *S* is a random set such that Pr(S = s) = p(s).

• The inclusion probability $\pi_k, k \in U$, can be derived from the sampling design

$$\pi_k = \sum_{s \ni k} p(s).$$

• Horvitz-Thompson estimator

$$\widehat{Y} = \sum_{k \in S} \frac{y_k}{\pi_k}.$$

• Unbiased if $\pi_k > 0$ for all $k \in U$.

Auxiliary information

| Auxiliary Information Auxiliary variables | Interest Information Interest variables |
|--|--|
| X | Y |
| known or partially known | unknown |





Regression estimator



Regression estimator $\widehat{\overline{Y}}_{\mathsf{REG}} = \widehat{\overline{Y}} + (\overline{X} - \widehat{\overline{X}})\widehat{b}$

Generalized regression estimator (GREG) (1)

- Multivariate auxiliary information given by the totals of p auxiliary variables x1,...,xp.
- Vector x_k = (x_{k1}, ..., x_{kj}, ..., x_{kp})' of values taken by the p auxiliary variables on unit k.
- The total $\mathbf{X} = \sum_{k \in U} \mathbf{x}_k$, is assumed to be known.
- The aim is to estimate $Y = \sum_{k \in U} y_k$, using the information given by X.

Generalized regression estimator (GREG) (2)

• GREG estimator:
$$\widehat{Y}_{GREG} = \widehat{Y} + (\mathbf{X} - \widehat{\mathbf{X}})'\widehat{\mathbf{b}}$$

• $\mathbf{X} = \sum_{k \in U} \mathbf{x}_k$

•
$$\widehat{\mathbf{X}} = \sum_{k \in S} \frac{\mathbf{x}_k}{\pi_k}$$

•
$$\widehat{\mathbf{b}} = \left(\sum_{k \in S} \frac{q_k \mathbf{x}_k \mathbf{x}'_k}{\pi_k}\right)^{-1} \sum_{k \in S} \frac{q_k \mathbf{x}_k y_k}{\pi_k}$$

• The q_k are weights.

$$\begin{split} \widehat{Y}_{\mathsf{GREG}} &= \sum_{k \in S} w_k y_k = \sum_{k \in S} \frac{g_k y_k}{\pi_k}, \\ w_k &= \frac{1}{\pi_k} \left\{ 1 + (\mathbf{X} - \widehat{\mathbf{X}})' \widehat{\mathbf{T}}^{-1} q_k \mathbf{x}_k \right\}, \\ g_k &= 1 + (\mathbf{X} - \widehat{\mathbf{X}})' \widehat{\mathbf{T}}^{-1} q_k \mathbf{x}_k \ , \\ \widehat{\mathbf{T}} &= \sum_{k \in S} \frac{\mathbf{x}_k \mathbf{x}'_k q_k}{\pi_k}. \end{split}$$

PROBLEM: the weights can be negative.

Ratio estimator 1

If the regression line crosses the origin.



• Let $\mathbf{x}_k = x_k$ only one auxiliary variable, and $q_k = \frac{1}{x_k}$. Then

•
$$\widehat{\mathbf{b}} = \left(\sum_{k \in S} \frac{q_k \mathbf{x}_k \mathbf{x}'_k}{\pi_k}\right)^{-1} \sum_{k \in S} \frac{q_k \mathbf{x}_k y_k}{\pi_k}$$

= $\left(\sum_{k \in S} \frac{x_k}{\pi_k}\right)^{-1} \sum_{k \in S} \frac{y_k}{\pi_k} = \frac{\widehat{Y}}{\widehat{X}}$

The regression estimator becomes the Ratio estimator

$$\widehat{Y}_{\mathsf{GREG}} = \widehat{Y} + (\mathbf{X} - \widehat{\mathbf{X}})'\widehat{\mathbf{b}} = \widehat{Y} + (X - \widehat{X})\frac{\widehat{Y}}{\widehat{X}} = \widehat{Y}\frac{X}{\widehat{X}}$$

- Multivariate auxiliary information given by the totals of p auxiliary variables x1,...,xp.
- Vector x_k = (x_{k1}, ..., x_{kj}, ..., x_{kp})' of values taken by the p auxiliary variables on unit k.
- The total

$$\mathbf{X} = \sum_{k \in U} \mathbf{x}_{\mathbf{k}},$$

is assumed to be known.

• The aim is to estimate $Y = \sum_{k \in U} y_k$, using the

information given by X.

• Horvitz-Thompson estimator $\sum_{k \in S} d_k y_k$, where $d_k = 1/\pi_k$.

• The idea consists of looking for new weights w_k as close as possible to d_k and such that

$$\sum_{k \in S} w_k \mathbf{x}_k = \mathbf{X} \text{ (calibration constraint)}.$$

• A pseudo-distance $G_k(.,.)$ between w_k and $d_k = 1/\pi_k$ is minimized,

$$\min_{w_k} \sum_{k \in S} \frac{G_k(w_k, d_k)}{q_k},$$

under the constraints of calibration.

- $q_k, k \in S$, are strictly positive known coefficients.
- Function $G_k(.,.)$ is assumed to be strictly convex, positive and such that $G_k(d_k, d_k) = 0$.

• The weights w_k are then defined by

$$w_k = d_k F_k(\boldsymbol{\lambda}' \mathbf{x}_k),$$

where $d_k F_k(.)$ is the reciprocal of the function $G'_k(., d_k)/q_k$, with

$$G'_k(w_k, d_k) = \frac{\partial G_k(w_k, d_k)}{\partial w_k},$$

and λ is the Lagrange multiplier following from the constraints.

- The vector $oldsymbol{\lambda}$ is obtained by solving the calibration equations:

$$\sum_{k \in S} w_k \mathbf{x}_k = \sum_{k \in S} d_k F_k(\boldsymbol{\lambda}' \mathbf{x}_k) \mathbf{x}_k = \sum_{k \in U} \mathbf{x}_k.$$

This system of equations can be non-linear (use of Newton method).

- Next the weights are computed $w_k = d_k F_k(\boldsymbol{\lambda}' \mathbf{x}_k)$.
- Finally, the calibrated estimator is $\widehat{Y}_{CAL} = \sum_{k \in S} w_k y_k$

- The weights w_k are close to $d_k = 1/\pi_k$ and are such that $\sum_{k \in S} w_k \mathbf{x}_k = \sum_{k \in U} \mathbf{x}_k$
- g-weights $g_k = \pi_k w_k$ (close to 1).
- Horvitz-Thompson estimator $\widehat{Y} = \sum_{k \in S} \frac{y_k}{\pi_k}$
- Calibrated estimator $\widehat{Y}_C = \sum_{k \in S} w_k y_k = \sum_{k \in S} \frac{g_k y_k}{\pi_k}.$

The g_k are the distorsion of the weights with respect to the Horvitz-Thompson estimator.

Suppose that $G_k(.,.)$ is chi-square function,

$$G_k(w_k, d_k) = \frac{(w_k - d_k)^2}{d_k},$$



Chi-square distance (2)

The derivative is

$$G'_k(w_k, d_k) = \frac{2(w_k - d_k)}{d_k},$$



Chi-square distance (3)

Calibration function





• Weights
$$w_k = d_k F_k(u) = d_k (1 + q_k \boldsymbol{\lambda}' \mathbf{x}_k).$$

The calibration equation is linear

$$\mathbf{X} = \widehat{\mathbf{X}} + \sum_{k \in S} d_k \mathbf{x}_k q_k \mathbf{x}'_k \boldsymbol{\lambda}$$

• Identification of $oldsymbol{\lambda}$

$$\boldsymbol{\lambda} = \left(\sum_{k \in S} d_k \mathbf{x}_k q_k \mathbf{x}'_k\right)^{-1} (\mathbf{X} - \widehat{\mathbf{X}})$$

Chi-square distance (5)

Weights

$$w_{k} = d_{k}F_{k}(u) = d_{k}(1 + q_{k}\boldsymbol{\lambda}'\mathbf{x}_{k})$$
$$= d_{k}\left[1 + q_{k}(\mathbf{X} - \widehat{\mathbf{X}})'\left(\sum_{k \in S} \frac{\mathbf{x}_{k}q_{k}\mathbf{x}'_{k}}{\pi_{k}}\right)^{-1}\mathbf{x}_{k}\right]$$

The calibrated estimator is then equal to the generalized regression estimator which is

$$\widehat{Y}_{\text{GREG}} = \widehat{Y} + (\mathbf{X} - \widehat{\mathbf{X}})'\widehat{\mathbf{b}},$$

where

$$\widehat{\mathbf{b}} = \widehat{\mathbf{T}}^{-1} \sum_{k \in S} \frac{\mathbf{x}_k y_k q_k}{\pi_k}.$$

and

$$\widehat{\mathbf{T}} = \left(\sum_{k \in S} \frac{\mathbf{x}_k \mathbf{x}'_k q_k}{\pi_k}\right)$$

The raking ratio method: distance

Suppose that the distance is

$$G(w_k, d_k) = w_k \log \frac{w_k}{d_k} + d_k - w_k.$$



"Raking ratio": function $G(w_k, = d_k)$ with $q_k = 1$ and $d_k = 10$

The raking ratio method: derivatives

The derivative of the distance is

$$G'(w_k, d_k) = \log \frac{w_k}{d_k},$$



The raking ratio method: calibration

The calibration function is

 $F_k(u) = \exp q_k u.$



"Raking ratio": function $F_k(u)$ with $q_k = 1$ ADVANTAGE: The weights are always positive.

Marginal calibration(1)

Adjust the following table

| 80 | 170 | 150 | 400 |
|-----|-----|-----|------|
| 90 | 80 | 210 | 380 |
| 10 | 80 | 130 | 220 |
| 180 | 330 | 490 | 1000 |

to the marginal column (430, 360, 210), and the marginal row (150, 300, 550).

| Calibration by row: iteration 1 | | | | |
|---------------------------------|--------|--------|---------|--|
| 86.00 | 161.25 | 430.00 | | |
| 85.26 | 75.79 | 198.95 | 360.00 | |
| 9.55 | 76.36 | 124.09 | 210.00 | |
| 180.81 | 334.90 | 484.29 | 1000.00 | |

| Calibration by column: iteration 2 | | | | | |
|------------------------------------|-------------|--------|---------|--|--|
| 71.35 | 163.70 | 183.13 | 418.18 | | |
| 70.73 | 70.73 67.89 | | 364.57 | | |
| 7.92 | 7.92 68.41 | | 217.25 | | |
| 150.00 | 300.00 | 550.00 | 1000.00 | | |
| Calibration by row: iteration 3 | | | | | |
| 73.36 | 168.33 | 188.31 | 430.00 | | |
| 69.85 67.04 | | 223.11 | 360.00 | | |
| 7.65 66.12 | | 136.22 | 210.00 | | |
| 150.87 | 201 /0 | 517 61 | 1000 00 | | |

| Calibration by column: iteration 4 | | | | | |
|------------------------------------|-------------|--------|---------|--|--|
| 72.94 | 167.50 | 189.12 | 429.56 | | |
| 69.45 | 69.45 66.71 | | 360.23 | | |
| 7.61 | 7.61 65.79 | | 210.22 | | |
| 150.00 | 300.00 | 550.00 | 1000.00 | | |
| Calibration by row: iteration 5 | | | | | |
| 73.02 | 167.67 | 189.31 | 430.00 | | |
| 69.40 66.67 | | 223.93 | 360.00 | | |
| 7.60 65.73 | | 136.67 | 210.00 | | |
| 150.02 | 300.06 | 549.91 | 1000.00 | | |

| Calibration by column: iteration 6 | | | | | | |
|------------------------------------|--------------------|--------|-----------|--|--|--|
| 73.01 | 73.01 167.64 189.3 | | 429.98 | | | |
| 69.39 | 69.39 66.65 223.97 | | 7 360.01 | | | |
| 7.60 | 65.71 | 136.69 | 210.01 | | | |
| 150.00 | 300.00 | 550.00 |) 1000.00 | | | |
| Calibration by row: iteration 7 | | | | | | |
| 73.01 | 167.64 | 189.35 | 430.00 | | | |
| 69.39 | 66.65 | 223.96 | 360.00 | | | |
| 7.60 | 65.71 | 136.69 | 210.00 | | | |
| 150.00 | 300.00 | 550.00 | 1000.00 | | | |

| Calibration by column: iteration 8 | | | | | |
|------------------------------------|---------------------|--------|---------|--|--|
| 73.01 | 73.01 167.64 189.35 | | | | |
| 69.39 | 66.65 | 223.96 | 360.00 | | |
| 7.60 | 65.71 | 136.69 | 210.00 | | |
| 150.00 | 300.00 | 550.00 | 1000.00 | | |

After 8 iterations, the adjustment is very accurate.

Pseudo-distances

| lpha | $G^{lpha}(w_k,d_k)$ | $g^lpha(w_k,d_k)$ | $F_k^{lpha}(u)$ | Type |
|------|--|--|-----------------------|--------------------|
| 2 | $\frac{(w_k - d_k)^2}{2d_k}$ | $rac{w_k}{d_k}-1$ | $1 + q_k u$ | Chi-square |
| 1 | $w_k \log \frac{w_k}{d_k} + d_k - w_k$ | $\log \frac{w_k}{d_k}$ | $\exp(q_k u)$ | Entropy |
| 1/2 | $2(\sqrt{w_k} - \sqrt{d_k})^2$ | $2\left(1-\sqrt{rac{d_k}{w_k}} ight)$ | $(1 - q_k u/2)^{-2}$ | Hellinger Distance |
| 0 | $d_k \log \frac{d_k}{w_k} + w_k - d_k$ | $1 - \frac{d_k}{w_k}$ | $(1-q_k u)^{-1}$ | Inverse Entropy |
| -1 | $\frac{(w_k - d_k)^2}{2w_k}$ | $\left(1 - \frac{d_k^2}{w_k^2}\right)/2$ | $(1 - 2q_k u)^{-1/2}$ | Inverse Chi-square |

Logistic function: distance



Figure 1: function $G(w_k, = d_k)$ with $q_k = 1$ and $d_k = 10$

The derivative of the distance



function $G'(w_k, d_k)$ with $q_k = 1$ and $d_k = 10$

Logistic function



The weights can be bounded in such a way that

$$B^- \le \frac{w_k}{d_k} \le B^+.$$

For instance $B^- = 0.4$ and $B^+ = 3$.

- Other calibration functions can also be used.
- The variance of the regression estimator is a variance of residuals.

Remarks on the calibration problem 1

- Calibration to several stages (municipalities, households, individuals)
- If the calibration variables can explain the nonresponse, then a calibration can be used to correct at the same time the sampling error and the nonresponse error.

A good software of calibration?

- Easy to use?
- Which distance can be used?
- Possibility to impose bound?
- Special functionalities for non-responses?
- Computation of the estimator of variance, or at least of the residuals?
- Shareware?

References

Deville, J.-C. and Särndal, C.-E. (1992). Calibration estimators in survey sampling. *Journal of the American Statistical Association*, 87:376–382.

Deville, J.-C., Särndal, C.-E., and Sautory, O. (1993). Generalized raking procedure in survey sampling. *Journal of the American Statistical Association*, 88:1013–1020.

Tillé, Y. (2001). *Théorie des sondages: échantillonnage et estimation en populations finies*. Dunod, Paris.